



Antinuclear antibodies and subclinical interstitial lung disease in community-dwelling adults: the MESA study

To the Editor:

The presence of a systemic autoimmune rheumatic disease (ARD) is a well-known risk factor for interstitial lung disease (ILD). For example, 33% of adults with rheumatoid arthritis (RA) have subclinical ILD [1]. Higher serum levels of IgM rheumatoid factor (RF), IgA RF, and anti-cyclic citrullinated peptide antibody 2 are associated with subclinical ILD in community-dwelling adults [2]. It is unknown whether this relationship between autoimmunity and subclinical ILD is limited to RA-related autoantibodies, or extends more broadly to other epitopes. High attenuation areas (HAA) and interstitial lung abnormalities (ILA) are validated quantitative and qualitative subclinical ILD phenotypes, respectively. In community-dwelling adults, greater HAA is associated with reduced forced vital capacity, reduced exercise capacity, elevated serum levels of matrix metalloproteinase-7 and interleukin-6, higher prevalence of ILA on computed tomography (CT) scans of the chest, higher all-cause mortality rate, and an increased risk of developing clinically evident ILD and ILD-specific mortality at 12-year follow-up [3, 4]. ILA has been associated with all-cause mortality in four different longitudinal cohorts [5]. The purpose of this study was to examine the association between antinuclear antibody (ANA) and both HAA and ILA in community-dwelling adults enrolled in the Multi-Ethnic Study of Atherosclerosis (MESA).

MESA is a population-based cohort study of 6814 adults aged 45–84 years when they were enrolled from 2000 to 2002 without regard to lung disease or ARD [6]. Cardiac CT scans were performed in 6812 participants at examination 1 (2000–2002) [7] and full lung CT scans in 2907 participants at examination 5 (2010–2012). Measurement of HAA, defined as the percentage of lung volume with attenuation values between –600 and –250 Hounsfield units, and ILA, defined as ground glass abnormalities, reticular abnormalities, diffuse centrilobular nodularity, nonemphysematous cysts, honeycombing, and traction bronchiectasis affecting >5% of a lung zone in a non-dependent manner [8], has been previously described [3]. HAA was quantified on the 6812 examination 1 cardiac CT scans. Each of the 2907 examination 5 full lung CT scans was visually inspected by one expert radiologist for the presence or absence of ILA. ANA was measured in frozen examination 1 sera from 6626 participants using indirect immunofluorescence with HEp-2 cell substrate at TheraTest Labs (TheraTest Labs Inc, Lombard, IL, USA) [9]. Intra-assay coefficient of variation was <10% [9]. ANA level was expressed in units. An ANA value >10 units was defined as positive. We examined the linearity of the associations between ANA and both HAA and ILA using generalised additive models with loess smoothing functions. We used multiple linear regression to examine associations between natural log-transformed ANA and natural log-transformed HAA, controlling for age, sex, race/ethnicity, body mass index (BMI), height, waist circumference, pack-years of smoking, current smoking status, estimated glomerular filtration rate, study site, education, total imaged lung volume, percent emphysema, and tube current. To ease interpretation of our beta coefficients of natural log-transformed ANA, we have presented base 2 exponentiated beta coefficients, which are the percent differences in HAA per doubling of ANA. We estimated prevalence ratios (PR) for the associations between log₂-transformed ANA and ILA using Poisson regression with robust standard error estimation, controlling for age, sex, race/ethnicity, pack-years of smoking, and current smoking status. We performed analyses stratified by age, sex, race/ethnicity, smoking status and BMI. We used likelihood ratio



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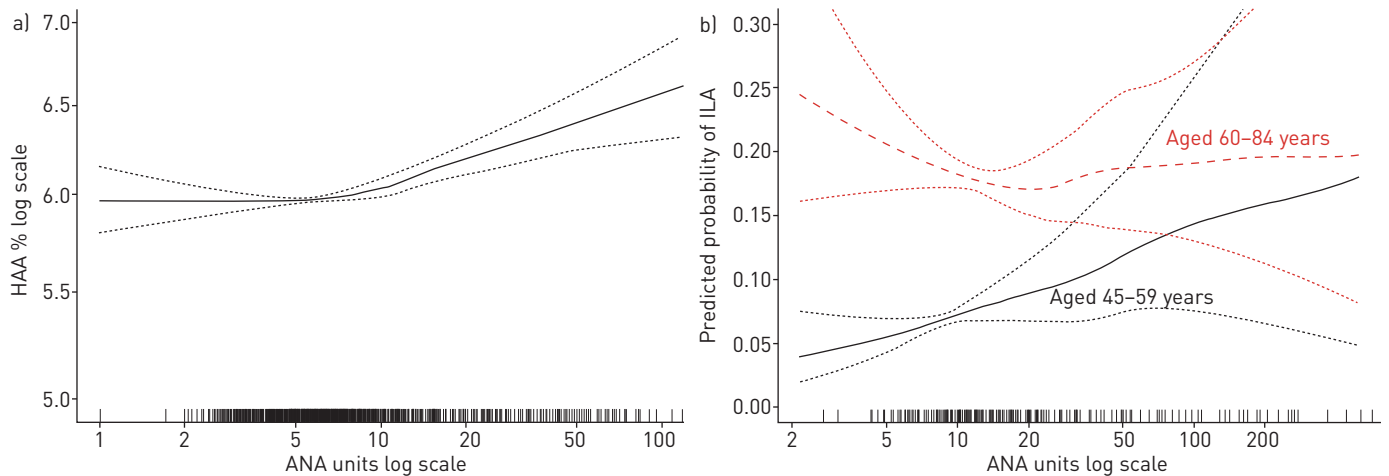


FIGURE 1 a) Continuous relationship of antinuclear antibody at examination 1 with predicted (adjusted) percent high attenuation areas at examination 1 ($n=6626$). Smoothed regression line (solid black line) is adjusted for age, sex, race/ethnicity, body mass index, height, waist circumference, pack-years of smoking, current smoking status, estimated glomerular filtration rate, study site, education, total imaged lung volume, percent emphysema, and tube current. Dotted lines are the 95% confidence bands. Each vertical tick mark on the rug plot along the internal border of the x-axis represents one study participant. Overall p-value for association <0.001 , p-value for non-linearity 0.002. b) Continuous relationship of antinuclear antibody at examination 1 with the predicted probability of interstitial lung abnormalities at examination 5 ($n=2366$). Smoothed regression lines (black solid line: age 45–59 years at examination 1; red dashed line: age 60–84 years at examination 1) are adjusted for age, sex, race/ethnicity, pack-years of smoking, and current smoking status. Dotted lines are the 95% confidence bands. Each vertical tick mark on the rug plot along the internal border of the x-axis represents one study participant. In participants aged 45–59 years ($n=1256$), overall p-value for association 0.009, p-value for non-linearity 0.58. In participants aged 60–84 years ($n=1110$), overall p-value for association 0.68, p-value for non-linearity 0.12.

tests to test for effect modification, and multiple imputation by chained equations to account for missing covariate data [10]. Only 0.4% of participants had any missing data. Analyses were performed in STATA, version 15.1 (College Station, TX, USA) and R, version 3.6.1 (R Foundation for Statistical Computing, Vienna, Austria).

The baseline characteristics of the MESA cohort have been previously published [3]. Of the 2430 participants with non-equivocal measurements of ILA at examination 5, ANA was measured in 2366 at examination 1. Thus, 6626 and 2366 participants were included in the HAA and ILA analyses, respectively. Mean age at examination 1 was 62 ± 10 years; 53% (3516/6626) were female. Of the 6626 participants included in the HAA analyses, 39% identified themselves as white, 27% as African American, 12% as Chinese American, and 22% as Hispanic. 41% (2688/6623) were former smokers; 14% (927/6623) were current smokers. 11% (741/6626) were ANA positive. Median ANA was 4 units (interquartile range (IQR) 3–7). Median HAA was 5.62% (IQR 4.55–7.19%). ILA prevalence was 12.4% (293/2366).

In an unadjusted model, HAA at examination 1 increased by 3.50% (95% CI 2.25 to 4.77%; $p<0.001$) per doubling of ANA at examination 1. In a fully adjusted model, HAA increased by 1.83% (95% CI 1.12 to 2.55%; $p<0.001$) per doubling of ANA (figure 1a). In a fully adjusted model, the p-value for the interaction between ANA and race/ethnicity was 0.04. In fully adjusted models, HAA increased by 2.82% (95% CI 1.49 to 4.16%; $p<0.001$) among African Americans, 2.93% (95% CI 0.85 to 5.06%; $p=0.006$) among Chinese Americans, and 2.14% (95% CI 0.65 to 3.66%; $p=0.005$) among Hispanic subjects per doubling of ANA. There was no statistically significant association between ANA and HAA among white subjects (0.65% increase per doubling of ANA, 95% CI -0.46 to 1.76%; $p=0.25$). There was no statistically significant interaction between ANA and age, sex, smoking status or BMI (p-value for interaction >0.50 for each).

In an unadjusted model, the prevalence of ILA at examination 5 increased by 17% per doubling of ANA at examination 1 (PR 1.17, 95% CI 1.04–1.33; $p=0.01$). However, in a fully adjusted model, there was no statistically significant association between ANA and ILA prevalence (PR 1.07, 95% CI 0.94–1.22; $p=0.29$). In a fully adjusted model, the p-value for the interaction between ANA and age was 0.003. In a fully adjusted model, ILA increased by 33% per doubling of ANA (PR 1.33, 95% CI 1.09–1.63; $p=0.006$) among younger participants (age 45–59 years at examination 1) (figure 1b). There was no statistically significant association between ANA and ILA prevalence among older participants (age 60–84 years at examination 1) (PR 0.97, 95% CI 0.83–1.13, p-value=0.68). There was no statistically significant interaction between ANA and sex, race/ethnicity, or smoking status (p-value for interaction >0.10 for each).


In this study, we demonstrated a positive association between levels of ANA, a marker of autoimmunity, and HAA, a quantitative CT biomarker of subclinical ILD. The association between ANA and HAA was strongest among non-white participants and the association between ANA and ILA was stronger among younger participants. Our results indicate that the relationship between autoimmunity and subclinical ILD is not limited to RA-related autoantibodies, but includes other epitopes as well.

It is well-established that autoantibodies precede the development of clinical manifestations of ARDs by several years [11, 12]. ANA was detected in the serum of 77% (89/115) of military personnel with systemic lupus erythematosus (SLE) [11] and in the serum of 44% (23/44) of Swedish patients with Sjogren's syndrome prior to symptom onset [12]. Our findings suggest that higher serum ANA levels may also be a risk factor for subclinical ILD.

It is perhaps not surprising that race/ethnicity and age modified the effect of ANA on HAA and ILA, respectively. The prevalence of ANA is modestly higher among African Americans than among non-Hispanic whites in the USA [13], and the prevalence of certain ARDs is higher in non-white populations. For example, African Americans, Asian Americans, and Hispanic Americans have a higher prevalence of SLE than do white subjects in the USA [14]. Moreover, non-white populations tend to have more severe manifestations of ARDs. Among patients with systemic sclerosis, African Americans have a higher prevalence of ILD and more severe ILD than do whites [15]. Although the prevalence of ANA increases with age [13], ARDs such as SLE often present at younger ages [14].

Our study has some limitations. It was cross-sectional, which limits our ability to make causal inferences. Because it was observational, our results may be confounded by unmeasured or poorly measured potential confounders. However, data in MESA were measured with great precision, and we controlled for potential confounders using multivariate modelling approaches to minimise residual confounding.

In summary, our findings provide additional support for the relationship between autoimmunity and subclinical ILD. They suggest autoimmunity may play a role in the pathogenesis of subclinical ILD, even among individuals without an established ARD. Future studies should characterise the specific autoantibody profiles associated with subclinical ILD.

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